

Definition of the Probability Characteristic of the System from Given Region for Case $(2^+, 1^-)$

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Abstract— In this paper has been studied the process of semi-markovian random walk with jumps and two delaying screens. The Laplace transformation of the distribution of a random variable $\tau(\omega)$ is obtained.

Keywords— process of semi-markovian random walk; the probability space; Laplas transformation

I. INTRODUCTION

Investigation of the ergodic distribution for semi-markovian random walk take process a special place in the theory of random processes. In 1975 V.Smit has proved the ergodic theorem for semi-markovian processes [1]. The general theorem about ergodic for processes with discrete intervention is proved in [2]. In [3] the ergodic theorem for complex semi-markovian processes with delaying screen is proved.

In [4] to find the Laplace transformation of the distribution for case $(1^+, 1^-)$ of a random variable $\tau(\omega)$.

II. THE PROCESS CONSTRUCTION

Let the sequence $\{\xi_k, \eta_k\}_{k=1, \infty}$, where $\xi_k, \eta_k, k=1, \infty$, are independent identically distributed random variables and independent themselves, $\xi_k > 0$ is given on the probability space $(\Omega, \mathfrak{F}, P(\cdot))$.

We construct the process [5]

$$X_1(t) = \sum_{k=0}^{m-1} \eta_k, \text{ if } \sum_{k=0}^{m-1} \xi_k \leq t < \sum_{k=0}^m \xi_k, \quad m = \overline{1, \infty}$$

where $\eta_0 = z \geq 0, \xi_0 = 0$.

We delay process $X_1(t)$ with screen in the zero (see [1]):

$$X_2(t) = X_1(t) - \inf_{0 \leq s \leq t} (0, X_1(s))$$

Then we delay this process with screen in $a(a > 0)$:

$$X(t) = X_2(t) - \sup_{0 \leq s \leq t} (0, X_2(s) - a).$$

This process is called the process of semimarkov random walk with double delaying screens in the "a" end zero.

We introduce a random variable $\tau(\omega)$, meaning the duration of the time in which process $X(t)$ is in a region $(0, a)$.

III. THE FINDING OF THE LAPLACE TRANSFORMATION OF THE DISTRIBUTION OF A RANDOM VARIABLE $\tau(\omega)$

The purpose in this paper to find the Laplace transformation of the distribution of a random variable $\tau(\omega)$. We denote

$$K(t|z) = P\{\tau > t | X(0) = z\}.$$

It is obvious, that

$$K(t|z) = P\left\{\inf_{0 \leq s \leq t} X(s) > 0; \sup_{0 \leq s \leq t} X(s) < a | X(0) = z\right\}$$

On total probability form we have:

$$\begin{aligned} K(t|z) &= P\left\{\inf_{0 \leq s \leq t} X(s) > 0; \sup_{0 \leq s \leq t} X(s) < a; \xi_1 > t | X(0) = z\right\} + \\ &+ P\left\{\inf_{0 \leq s \leq t} X(s) > 0; \sup_{0 \leq s \leq t} X(s) < a; \xi_1 < t | X(0) = z\right\} = \\ &= P\{\xi_1(\omega) > t\} + \\ &+ \int_{s=0}^t \int_{y=0}^a P\{\xi_1(\omega) \in ds; z + \eta_1 \in dy\} P\{\tau > t - s | X(0) = y\} \end{aligned} \quad (1)$$

Then the equation (1) will be written in the following form:

$$\begin{aligned} K(t|z) &= P\{\xi_1(\omega) > t\} + \\ &+ \int_{s=0}^t \int_{y=0}^a P\{\xi_1(\omega) \in ds\} d_y P\{\eta_1 < y - z\} K(t - s|y) \end{aligned} \quad (2)$$

Let's denote:

$$\tilde{K}(\theta|z) = \int_{t=0}^{\infty} e^{-\theta t} K(t|z) dt, \quad \theta > 0. \quad (3)$$

$$\varphi(\theta) = Ee^{-\theta \xi_1}, \quad \theta > 0.$$

If to apply the Laplace transformation on both sides of the equation (1) with respect to t :

$$\int_{t=0}^{\infty} e^{-\theta t} \tilde{K}(t|z) dt = \int_{t=0}^{\infty} e^{-\theta t} P\{\xi_{S_1}(\omega) > t\} dt +$$

$$+ \int_{y=0}^a d_y P\{\eta_1 < y - z\} \int_{t=0}^{\infty} e^{-\theta t} \int_{s=0}^t P\{\xi_{S_1}(\omega) \in ds\} K(t-s|y) dy =$$

$$= \frac{1 - \varphi(\theta)}{\theta} + \varphi(\theta) \int_{y=0}^a \tilde{K}(\theta|y) d_y P\{\eta_1 < y - z\},$$

then we have the following equation for $\tilde{K}(\theta|z)$:

$$\tilde{K}(\theta|z) = \frac{1 - \varphi(\theta)}{\theta} + \varphi(\theta) \int_{y=0}^a \tilde{K}(\theta|y) d_y P\{\eta_1 < y - z\}. \quad (4)$$

Let's solve this equation in the class for the Laplace distributions. For example, let

$$\eta_1 = \eta_1^+ + \eta_2^+ - \eta_1^-,$$

$$F\{\eta_1 < t\} = \begin{cases} \frac{\lambda^2}{(\lambda + \mu)^2} e^{\mu t}, & t < 0, \\ 1 - \frac{\mu}{\lambda + \mu} \left[1 + \frac{\lambda}{\lambda + \mu} + \lambda t \right] e^{-\lambda t}, & t > 0. \end{cases} \quad (5)$$

Hence we have

$$P_{\eta_1}(t) = \begin{cases} \frac{\lambda^2 \mu}{(\lambda + \mu)^2} e^{\mu t}, & t < 0, \\ \frac{\lambda^2 \mu}{\lambda + \mu} \left[\frac{1}{\lambda + \mu} + t \right] e^{-\lambda t}, & t > 0. \end{cases} \quad (6)$$

and

$$\tilde{K}(\theta|z) = \frac{1 - \varphi(\theta)}{\theta} + \varphi(\theta) \frac{\lambda^2 \mu}{(\lambda + \mu)^2} e^{-\mu z} \int_{y=0}^z e^{\mu y} \tilde{K}(\theta|y) e^{\mu y} dy +$$

$$+ \varphi(\theta) \frac{\lambda^2 \mu}{(\lambda + \mu)^2} e^{\lambda z} \int_{y=z}^a e^{-\lambda y} \tilde{K}(\theta|y) dy -$$

$$- \varphi(\theta) \frac{\lambda^2 \mu z}{\lambda + \mu} e^{-\lambda z} \int_{y=0}^z e^{-\lambda y} \tilde{K}(\theta|y) dy +$$

$$+ \varphi(\theta) \frac{\lambda^2 \mu}{\lambda + \mu} e^{\lambda z} \int_{y=z}^a e^{-\lambda y} y \tilde{K}(\theta|y) dy. \quad (7)$$

We denote:

$$\int_0^{\infty} P\{\tau > t | X(0) = z\} dt = E(\tau | X(0) = z)$$

$$\tilde{K}(\theta|z) = \int_0^{\infty} e^{-\theta t} P\{\tau > t | X(0) = z\} dt,$$

$$L(\theta|z) = 1 - \theta \tilde{K}(\theta|z). \quad (8)$$

We can write, the equation (8) in the following form using (7):

$$L(\theta|z) = \varphi(\theta) - \varphi(\theta) \frac{\lambda^2 \mu}{(\lambda + \mu)^2} e^{-\mu z} \int_{y=0}^z e^{\mu y} dy +$$

$$+ \varphi(\theta) \frac{\lambda^2 \mu}{(\lambda + \mu)^2} e^{-\mu z} \int_{y=0}^z e^{\mu y} L(\theta|y) e^{\mu y} dy -$$

$$- \varphi(\theta) \frac{\lambda^2 \mu}{(\lambda + \mu)^2} e^{\lambda z} \int_{y=0}^z e^{-\lambda y} dy +$$

$$+ \varphi(\theta) \frac{\lambda^2 \mu}{(\lambda + \mu)^2} e^{\lambda z} \int_{y=z}^a e^{-\lambda y} L(\theta|y) dy + \varphi(\theta) \frac{\lambda^2 \mu z}{\lambda + \mu} e^{\lambda z} \int_{y=z}^a e^{-\lambda y} dy -$$

$$+ \varphi(\theta) \frac{\lambda^2 \mu z}{\lambda + \mu} e^{\lambda z} \int_{y=z}^a e^{-\lambda y} L(\theta|y) dy - \varphi(\theta) \frac{\lambda^2 \mu}{\lambda + \mu} e^{\lambda z} \int_{y=z}^a e^{-\lambda y} y dy +$$

$$+ \varphi(\theta) \frac{\lambda^2 \mu}{\lambda + \mu} e^{\lambda z} \int_{y=z}^a e^{-\lambda y} y L(\theta|y) dy. \quad (9)$$

From (9) we can receive the differential equation:

$$L'''(\theta|z) - (2\lambda - \mu)L''(\theta|z) + \lambda(\lambda - 2\mu)L'(\theta|z) +$$

$$+ \lambda^2 \mu [1 - \varphi(\theta)]L(\theta|z) = 0. \quad (10)$$

The characteristic equation of (10) will be in the following form

$$k^3(\theta) - (2\lambda - \mu)k^2(\theta) + \lambda(\lambda - 2\mu)k(\theta) + \lambda^2 \mu [1 - \varphi(\theta)] = 0. \quad (11)$$

Then the common solution of (10) will be

$$L(\theta|z) = \sum_{i=1}^3 d_i(\theta) e^{k_i(\theta)z}. \quad (12)$$

From (9) we can find the initial conditions for differential equation (10):

$$\left\{ \begin{aligned} L(\theta|a) &= \varphi(\theta) - \varphi(\theta) \frac{\lambda^2 \mu}{(\lambda + \mu)^2} e^{-\mu a} \int_{y=0}^a e^{\mu y} dy + \\ &+ \varphi(\theta) \frac{\lambda^2 \mu}{(\lambda + \mu)^2} e^{-\mu a} \int_{y=0}^a e^{\mu y} L(\theta|y) dy, \\ L'(\theta|a) &= \varphi(\theta) \frac{\lambda^2 \mu^2}{(\lambda + \mu)^2} e^{-\mu a} \int_{y=0}^a e^{\mu y} dy - \\ &- \varphi(\theta) \frac{\lambda^2 \mu^2}{(\lambda + \mu)^2} e^{-\mu a} \int_{y=0}^a e^{\mu y} L(\theta|y) dy, \\ L''(\theta|a) &= -\mu L'(\theta|a). \end{aligned} \right. \quad (13)$$

From (13) we can receive the following system of the linear algebraic equations for $d_1(\theta)$, $d_2(\theta)$ and $d_3(\theta)$.

$$\begin{aligned}
 & \left\{ \begin{aligned}
 & [(\lambda + \mu)^2 - (\mu + k_2(\theta))(\mu + k_3(\theta))]e^{k_1(\theta)a} + \\
 & (\mu + k_2)(\mu + k_3)e^{\mu a} \Big] d_1(\theta) + \\
 & [(\lambda + \mu)^2 - (\mu + k_1(\theta))(\mu + k_3(\theta))]e^{k_2(\theta)a} + \\
 & + (\mu + k_1)(\mu + k_3)e^{\mu a} \Big] d_2(\theta) + \\
 & [(\lambda + \mu)^2 - (\mu + k_1(\theta))(\mu + k_2(\theta))]e^{k_3(\theta)a} + \\
 & + (\mu + k_1(\theta))(\mu + k_2(\theta))e^{\mu a} \Big] d_3(\theta) = \\
 & = \varphi(\theta)(2\lambda\mu + \mu^2 - \lambda^2 e^{-\mu a}), \\
 & [(\lambda + \mu)^2 k_1(\theta) + \mu(\mu + k_2(\theta))(\mu + k_3(\theta))]e^{k_1(\theta)a} - \\
 & - \mu(\mu + k_2)(\mu + k_3)e^{\mu a} \Big] d_1(\theta) + \\
 & [(\lambda + \mu)^2 k_2(\theta) + \mu(\mu + k_1(\theta))(\mu + k_3(\theta))]e^{k_2(\theta)a} - \\
 & - \mu(\mu + k_1)(\mu + k_3)e^{\mu a} \Big] d_2(\theta) + \\
 & [(\lambda + \mu)^2 k_3(\theta) + \mu(\mu + k_1(\theta))(\mu + k_2(\theta))]e^{k_3(\theta)a} - \\
 & - \mu(\mu + k_1)(\mu + k_2)e^{\mu a} \Big] d_3(\theta) = \\
 & = \lambda^2 \mu \varphi(\theta)(1 - e^{-\mu a}), \\
 & (\mu + k_1(\theta))k_1(\theta)e^{k_1(\theta)a} + (\mu + k_2(\theta))k_2(\theta)e^{k_2(\theta)a} + \\
 & + (\mu + k_3(\theta))k_3(\theta)e^{k_3(\theta)a} = 0.
 \end{aligned} \right. \quad (14)
 \end{aligned}$$

To find $d_i(\theta), i = \overline{1,3}$ we must find $d_i(0), i = \overline{1,3}$.

It is obvious, that

$$\begin{aligned}
 L(\theta) &= \int_{z=0}^a L(\theta|z) dP\{\min(a, \eta_1^+) < z\} = \\
 &= \int_{z=0}^a L(\theta|z) d\left[1 - P\{\min(a, \eta_1^+) > z\}\right] = \\
 &= L(\theta|0)P\{\eta_1^+ > a\} - \int_{z=0}^a L(\theta|z) dz P\{\eta_1^+ > z\} \\
 L(\theta) &= L(\theta|a)(1 + \lambda a)e^{-\lambda a} + \lambda^2 \int_{z=0}^a ze^{-\lambda z} L(\theta|z) dz.
 \end{aligned} \quad (15)$$

For applications we find the expectation and variance of the distribution of the random variable $\tau(\omega)$. We know that

$$E\tau(\omega) = L'(0)$$

From (15) we find that

$$\begin{aligned}
 L'(0) &= -\frac{\lambda\mu a}{\lambda - 2\mu} \varphi'(0)e^{-\lambda a} - \frac{2\mu}{\lambda - 2\mu} \varphi'(0)e^{-\lambda a} + \frac{2\mu}{\lambda - 2\mu} \varphi'(0) + \frac{2\varphi'(0)}{f_1} \times \\
 &\times \left\{ \frac{1}{4(\lambda - 2\mu)} [(\lambda - \mu + b)(-\mu^3(2\lambda + \mu + b)^2 + 4\lambda^4\mu) + 2\lambda^5(2\lambda - \mu - b)] e^{\frac{2\lambda - 3\mu + b}{2} a} - \right. \\
 &- \frac{\lambda\mu(2\lambda - \mu + b)a}{8(\lambda - 2\mu)} [\mu^2(2\lambda + \mu + b)^2 - 4\lambda^4] e^{\frac{2\lambda - 3\mu + b}{2} a} + \\
 &+ \frac{1}{4(\lambda - 2\mu)} [(\lambda - \mu - b)(\mu^3(2\lambda + \mu - b)^2 - 4\lambda^4\mu) - 2\lambda^5(2\lambda - \mu - b)] e^{\frac{2\lambda - 3\mu - b}{2} a} + \\
 &+ \frac{\lambda\mu(2\lambda - \mu - b)a}{8(\lambda - 2\mu)} [\mu^2(2\lambda + \mu - b)^2 - 4\lambda^4] e^{\frac{2\lambda - 3\mu - b}{2} a} + \\
 &+ \left[\frac{\lambda^2 \mu^3(\lambda + \mu + b)}{\lambda - 2\mu} + \frac{2\lambda^3 \mu^4(\lambda + \mu + b)a}{(\lambda - 2\mu)(\mu - b)} \right] e^{-\frac{3\mu - b}{2} a} - \\
 &- \left[\frac{\lambda^2 \mu^3(\lambda + \mu - b)}{\lambda - 2\mu} + \frac{2\lambda^3 \mu^4(\lambda + \mu - b)a}{(\lambda - 2\mu)(\mu + b)} \right] e^{-\frac{3\mu + b}{2} a} - \\
 &- \left. \frac{\lambda\mu^3(2\lambda + \mu)b}{\lambda - 2\mu} e^{-\mu a} - \left[\frac{\lambda^3(\lambda^2 - 3\mu^2)b}{\lambda - 2\mu} + \lambda^4 \mu ab - \lambda^5 \mu a^2 b \right] e^{(\lambda - 2\mu)a} \right\}
 \end{aligned}$$

We know that

$$D\tau(\omega) = L''(0) - [L'(0)]^2.$$

The following fact is proved at $\lambda < 2\mu$:

$\lambda < 2\mu$	$E\tau(\omega)$
$a \rightarrow 0$	$-\varphi'(0) > 0$
$a \rightarrow \infty$	$\frac{2\mu}{\lambda - 2\mu} \varphi'(0) > 0$

REFERENCES

- [1] B.L. Smit, "The renewall processes", // Mathematics, 1961, vol. №36, p.5.
- [2] I.I. Gihman, A.V. Skorohod., "Theory of random processes", M., Nauka, 1973, vol.2, 639 p., in (Russian).
- [3] T.H. Nasirova "Complex process of semimarkovian random walk with screen", B.:Science, 1988, 50 p., in (Russian).
- [4] T.H. Nasirova, R.H. Sadiqova, "The Laplas transformation of the distribution of the length of the time of the stay in given region", AVT, ISSN 0132-4160, Riga, 2009, No.4, pp.30-36.,in (Russian).
- [5] A.A. Borovkov, "The stochastic process in the Queuing Theory", Moscow, Nauka, 1972, 368 p., in (Russian).